Conditional Independence in Continuous Domain¹

ETH zürich

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Based on joint work²³⁴ with:



Prof. Victor Panaretos, EPFL

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 $^{^2}$ Waghmare, K.G. and Panaretos, V.M., 2022. The Completion of Covariance Kernels.

³Waghmare, K.G. and Panaretos, V.M., 2024. Continuously Indexed Graphical Models.

⁴Waghmare, K.G. and Panaretos, V.M., 2023. The Positive-Definite Completion Problem.

Spine Bone Mineral Density: A Longitudinal Study⁵



■ Figure. The BMD measurements of ~100 individuals taken between the ages of 9 and 21 years.

⁵Bachrach, L.K., Hastie, T., Wang, M.C., Narasimhan, B. and Marcus, R., 1999. Bone mineral acquisition in healthy Asian, Hispanic, black, and Caucasian youth: a longitudinal study. *The Journal of Clinical Endocrinology & Metabolism*, 84 (12), pp.4702-4712. Image: Clinical Endocrinology & Metabolism, 84 (12), pp.4702-4712.

Spine Bone Mineral Density: A Longitudinal Study⁵



■ **Figure.** The BMD measurements of ~100 individuals taken between the ages of 9 and 21 years.

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• No individual was followed for longer than 5 years.

Covariance Estimation



• Figure. The pairs (s, t) s.t. BMD measurements of an individual for the both ages s and t are available.

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Covariance Estimation



- Figure. The pairs (s, t) s.t. BMD measurements of an individual for the both ages s and t are available.
- The covariance can be estimated only over a relatively small region around the diagonal.

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Covariance Recovery



• Let $X = \{X_t : t \in I\}$ be a 2nd-order stochastic process on an interval $I \subset \mathbb{R}$ with mean zero and covariance K.

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Covariance Recovery



- Let $X = \{X_t : t \in I\}$ be a 2nd-order stochastic process on an interval $I \subset \mathbb{R}$ with mean zero and covariance K.
- Given a consistent estimator \hat{K}_{Ω} of the partial covariance $K_{\Omega} = K|_{\Omega}$ is it possible to construct a consistent estimator of K?

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Covariance Recovery



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1. Under what conditions?

Covariance Recovery



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- Given a consistent estimator \hat{K}_{Ω} of the partial covariance $K_{\Omega} = K|_{\Omega}$ is it possible to construct a consistent estimator of K?

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- 1. Under what conditions?
- 2. How? And how accurate?

Covariance Completion



• Under what conditions is it possible to extend a kernel $K_{\Omega} : \Omega \to \mathbb{R}$ to a covariance kernel K on I?

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Covariance Completion



• Under what conditions is it possible to extend a kernel $K_{\Omega} : \Omega \to \mathbb{R}$ to a reproducing kernel K on I?

1. Is there a unique completion? If not, how many completions can there be?

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Covariance Completion



• Under what conditions is it possible to extend a kernel $K_{\Omega} : \Omega \to \mathbb{R}$ to a reproducing kernel K on I?

- 1. Is there a unique completion? If not, how many completions can there be?
- 2. Is there a special completion which has a nice interpretation in terms of the process X?

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Related Work

Fragmented Functional Data

- **Markov Chains:** Delaigle and Hall (2016).
- **2** Unique Completion: Delaigle, Hall, Huang, and Kneip (2021); Lin, Wang, and Zhong (2021), Descary and Panaretos (2019).

Positive-Definite Completion

- **1** Stationary on Z: Carathéodory (1907); Calderón and Pepinsky (1952).
- **2** Stationary on \mathbb{R} : Krein (1940); Rudin (1963).
- **3** Nonstationary on finite set: Dym and Gohberg (1981); Grone, Johnson, Sá, and Wolkowicz (1984).

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4 Nonstationary on \mathbb{R} : ?

Completion Formula



• Let K_{Ω} be a kernel such that

$$K_{I_1} = K_{\Omega}|_{I_1 \times I_1}$$
$$K_{I_2} = K_{\Omega}|_{I_2 \times I_2}$$

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are reproducing kernels.

Completion Formula



• Let K_{Ω} be a kernel such that

$$K_{I_1} = K_{\Omega}|_{I_1 \times I_1}$$
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are reproducing kernels.

• How to define K(s,t) such that K is a reproducing kernel?

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Completion Formula



• Consider the subkernel $K_{I_1 \cap I_2}$. It has an RKHS

$$\mathcal{H} = \mathcal{H}(K_{I_1 \cap I_2})$$

Completion Formula



• Consider the subkernel $K_{I_1 \cap I_2}$. It has an RKHS

 $\mathcal{H} = \mathcal{H}(K_{I_1 \cap I_2})$

• And consider the cross-covariances $K_{\Omega}(t, \cdot) : u \mapsto K_{\Omega}(t, u)$ $K_{\Omega}(\cdot, s) : u \mapsto K_{\Omega}(u, s)$

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for $u \in I_1 \cap I_2$.

Completion Formula



• Consider the subkernel $K_{I_1 \cap I_2}$. It has an RKHS

 $\mathcal{H} = \mathcal{H}(K_{I_1 \cap I_2})$

And consider the *cross-covariances*

$$K_{\Omega}(t, \cdot) : u \mapsto K_{\Omega}(t, u)$$
$$K_{\Omega}(\cdot, s) : u \mapsto K_{\Omega}(u, s)$$

for $u \in I_1 \cap I_2$.

Define

 $K(s,t) = \langle K_{\Omega}(s,\cdot), K_{\Omega}(\cdot,t) \rangle_{\mathcal{H}_{I_1 \cap I_2}}$ $(\Box) \langle \Box \rangle \langle$

Completion Formula



If we define $K_{\star}: I \times I \to \mathbb{R}$ as

$$K_{\star}(s,t) = \begin{cases} K_{\Omega}(s,t) & (s,t) \in \Omega\\ \langle K_{\Omega}(s,\cdot), K_{\Omega}(\cdot,t) \rangle_{\mathcal{H}} & \text{otherwise.} \end{cases}$$

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the result is a valid covariance on I.

Completion Algorithm



The procedure can be iterated in many ways, but regardless of the manner of completion one recovers the same completion.



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Estimation

Construction of K_{\star}



The formula (*) reduces to

$$\mathbf{R}_{p} = \left[\mathbf{J}_{p}^{-1/2}\mathbf{S}_{p}^{*}\right]^{*} \left[\mathbf{J}_{p}^{-1/2}\mathbf{D}_{p}\right]$$

Using the eigendecomposition

$$\mathbf{J}_p = \sum_{k=1}^\infty \lambda_{p,k} e_{p,k} \otimes e_{p,k}$$

we have

$$\boxed{\mathbf{R}_p = \sum_{k=1}^\infty \frac{1}{\lambda_{p,k}} \mathbf{S}_p e_{p,k} \otimes \mathbf{D}_p^* e_{p,k}}$$

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Estimation

Construction of \hat{K}_{\star}



Using the eigendecomposition

$$\hat{\mathbf{J}}_p = \sum_{k=1}^{\infty} \hat{\lambda}_{p,k} \hat{e}_{p,k} \otimes \hat{e}_{p,k}$$

we have

$$\hat{\mathbf{R}}_p = \sum_{k=1}^{N(p)} rac{1}{\hat{\lambda}_{p,k}} \hat{\mathbf{S}}_p \hat{e}_{p,k} \otimes \hat{\mathbf{D}}_p^* \hat{e}_{p,k}$$

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Illustration



Canonical Completion

Naturalness of Completion

If we have



$$K_{\Omega}(s,t) = \min(s,t)$$
 for $(s,t) \in \Omega$,

then the completion algorithm gives us

$$K_*(s,t) = \min(s,t)$$
 for $s, t \in I$.

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Canonical Completion

Naturalness of Completion

In fact, the same applies for

Analytic covariances:



$$K(s,t) = e^{-(s-t)^2}, \cos(s-t), \dots$$

• Covariances of Markov Gaussian process:

$$K(s,t) = \min(s,t), e^{-|s-t|}, \dots$$

■ *Nice* finite rank covariances:

$$K(s,t) = st + s^2 t^2, \dots$$



 6 Kneip, A. and Liebl, D., 2020. On the optimal reconstruction of partially observed functional data. The Annals of Statistics, 48(3), pp.1692-1717.

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Uniqueness of Completion



We define the Schur complement of a kernel K_I on I with respect to $J \subset I$ as the kernel K_I/K_J

$$= K_I(s,t) - \langle K_I(s,\cdot), K_I(\cdot,t) \rangle_{\mathcal{H}_J}$$

for $s, t \in I \setminus J$.

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Uniqueness of Completion



The partial covariance K_{Ω} admits a unique completion iff

$$K_{I_p}/K_{I_p \cap I_{p+1}} = 0$$
, for $1 \le p < r$
 $K_{I_{q+1}}/K_{I_q \cap I_{q+1}} = 0$, for $r \le q < m$.

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Consequently, for some I_r

$$X_t \in \operatorname{Span}\{X_u : u \in I_r\}$$

for every $t \in I$.

The Graph Ω



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The Graph Ω



• The region Ω can be thought of as a graph with the vertices $s \in I$ and edges $(s, t) \in \Omega$.

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The Graph Ω



- The region Ω can be thought of as a graph with the vertices $s \in I$ and edges $(s, t) \in \Omega$.
- It turns out that so long as Wseparates $s, t \in I$ with respect to Ω , we have

$$K_{\star}(s,t) = \langle K_{\star}(s,\cdot), K_{\star}(\cdot,t) \rangle_{\mathcal{H}_{W}}$$

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Conditional Independence



■ If X was a Gaussian process, then by *Loève isometry* this implies

$$\mathbb{E}\left[X_s X_t\right] = \mathbb{E}\left[\mathbb{E}\left[X_s | X_W\right] \mathbb{E}\left[X_t | X_W\right]\right]$$

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Conditional Independence



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And moreover,

$$\operatorname{Cov}(X_s, X_t | X_W) = 0$$

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Conditional Independence



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 $\mathbb{E}\left[X_s X_t\right] = \mathbb{E}\left[\mathbb{E}\left[X_s | X_W\right] \mathbb{E}\left[X_t | X_W\right]\right]$

And moreover,

$$\operatorname{Cov}(X_s, X_t | X_W) = 0$$

In other words,

$$X_s \perp\!\!\!\perp X_t \mid \{X_u : u \in W\}.$$

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Assume that U is an interval.

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Assume that U is an interval.

The region Ω can be interpreted as an "adjacency matrix". Two vertices $u, v \in U$ are adjacent if and only if $(u, v) \in \Omega$.

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Visualizing paths is tricky.



Visualizing paths is tricky. Consider for example the path

 v, x_1, x_2, x_3, u .

The unfilled circles \circ represent the edges (v, x_1) , (x_1, x_2) , (x_2, x_3) and (x_3, u) . For a valid path they must lie within Ω .

On the other hand * represents the non-edge (u, v).

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Separators are sets which intercept paths.

Here our path (v, x_1, x_2, x_3, u) passes through W because some of the points in the path lie in W. In fact, notice that this is true for all the paths from v to u in Ω .

Thus, W separates u and v in Ω .

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Continuously Indexed Graphical Model



We say that X has the graph $\Omega \subset U \times U$ if for every $u, v \in U$ separated by $W \subset U$ in Ω , it satisfies the global Markov property:

$$X_u \perp \!\!\!\perp X_v \mid X_W.$$

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Continuously Indexed Graphical Model



We say that X has the graph $\Omega \subset U \times U$ if for every $u, v \in U$ separated by $W \subset U$ in Ω , it satisfies the *separation equation*:

$$K(u,v) = \langle K(u,\cdot), K(\cdot,v) \rangle_{\mathcal{H}(K_{\mathbf{W}})}$$

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Continuously Indexed Graphical Model



We say that X has the graph $\Omega \subset U \times U$ if for every $u, v \in U$ separated by $W \subset U$ in Ω , it satisfies the *separation equation*:

 $K(u,v) = \langle K(u,\cdot), K(\cdot,v) \rangle_{\mathcal{H}(K_{\mathbf{W}})}$

A reproducing kernel characterization of conditional independence in a Gaussian process!

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Covariance Recovery





Covariance Recovery



• Assume that X admits a graphical structure with the graph $\tilde{\Omega}$, and we can estimate the covariance over $\bar{\Omega}$.

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Covariance Recovery



- Assume that X admits a graphical structure with the graph $\tilde{\Omega}$, and we can estimate the covariance over $\bar{\Omega}$.
- If $\tilde{\Omega} \subset \Omega \subset \bar{\Omega}$ then it is possible to recover the covariance K of X from $K_{\bar{\Omega}}$.

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More on Positive-Definite Completion

- 1. Using the canonical completion K_{\star} , one can characterize all completions of K_{Ω} .
- 2. One can derive more elegant expressions for K_{\star} :

$$K_{\star}(x,y) = -\frac{1}{2} [K_{\Omega}(x,x) + K_{\Omega}(y,y] + \sup_{f} \left[f(x) + f(y) - \frac{1}{2} \|f\|_{*}^{2} \right]$$

where

$$||f||_*^2 = ||f_{I_1}||^2 - ||f_{I_1 \cap I_2}||^2 + ||f_{I_2}||^2 - \dots + ||f_{I_p}||^2$$

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3. The concept of conditional independence supplies the correct definition of a special solution of positive-definite completion!

Recovering Graph from Covariance



Recovering Uncountably Large Graphs



We say that X has the graph $\Omega \subset U \times U$ if for every $u, v \in U$ separated by $W \subset U$ in Ω , it satisfies the *separation equation*:

$$K(u,v) = \langle K(u,\cdot), K(\cdot,v) \rangle_{\mathcal{H}(K_{\mathbf{W}})}$$

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This is not a very handy characterization unlike he multivariate case.

Resolving Uncountably Large Graphs

Let's partition the domain U into $\pi = \{U_1, \ldots, U_p\}.$



Instead of asking whether there is an edge between $u, v \in U$, we pose the related question of whether there is an edge between (some point in) U_i and (some point in) U_j .

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Resolving Uncountably Large Graphs

Turns out Ω^{π} admits a nice inverse zero characterization!

For $1 \leq i, j \leq p$, let $\mathbf{K}_{ij} : L^2(U_j, \mu) \to L^2(U_i, \mu)$ be the integral operator induced by the integral kernel $K_{ij} = K|_{U_i \times U_i}$ given by

$$\mathbf{K}_{ij}f(u) = \int_{U_j} K_{ij}(u, v) f(v) \ d\mu(v)$$

Define the *covariance operator matrix* \mathbf{K}_{π} induced by the partition π as $\mathbf{K}_{\pi} = [\mathbf{K}_{ij}]_{i,j=1}^{p}$. Furthermore, we define the correlation operator matrix \mathbf{R}_{π} induced by the partition π as $\mathbf{R}_{\pi} = [\mathbf{R}_{ij}]_{i,j=1}^{p}$ specified entrywise by

$$\mathbf{R}_{ij} = \mathbf{K}_{ii}^{-1/2} \mathbf{K}_{ij} \mathbf{K}_{jj}^{-1/2}.$$

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Resolving Uncountably Large Graphs

The correlation operator matrix \mathbf{R}_{π} behaves much better than \mathbf{K}_{π} . In fact, it is often invertible!

Theorem

Under some technical conditions, if \mathbf{R}_{π} is invertible, then the graph Ω^{π} is related to the inverse $\mathbf{P}_{\pi} = \mathbf{R}_{\pi}^{-1}$ as follows:

$$\Omega^{\pi} \equiv \lim_{\epsilon \to 0} (\Omega + \mathbb{B}_{\epsilon})^{\pi} = \bigcup \{ U_i \times U_j : \|\mathbf{P}_{ij}\| \neq 0 \}.$$

The \subset statement holds even without the said technical condition.

Choosing a finer partition π yields a higher resolution version Ω^{π} of Ω .

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Graphs of some processes



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Figure: The covariance K, precision matrix \mathbf{P}_{π} and the graph Ω^{π} of Gaussian kernel: $K(u, v) = \exp[-(u - v)^2]$.

Graphs of some processes



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Figure: The covariance K, precision matrix \mathbf{P}_{π} and the graph Ω^{π} of Brownian motion: $K(u, v) = \min(u, v)$.

Graphs of some processes



Figure: The covariance K, precision matrix \mathbf{P}_{π} and the graph Ω^{π} of integrated Brownian motion: $K(u, v) = \operatorname{Cov}(X_u, X_v)$ where X_t is given by $X_t = \int_{\max(0, t-1/2)}^t B_u \, du$.

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Publications

- **1.** The Completion of Covariance Kernels. Kartik G. Waghmare and Victor M. Panaretos. The Annals of Statistics, 50.6 (2022), pp. 3281–33.
- 2. Continuously Indexed Graphical Models. Kartik G. Waghmare and Victor M. Panaretos. Journal of the Royal Statistical Society: Series B (2024).
- 3. The Positive-Definite Completion Problem. Kartik G. Waghmare and Victor M. Panaretos. Transactions of the American Mathematical Society (2024).

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